

Theorem: (Cramer - Rao Inequality)

In any regular estimation case, the variance of an unbiased estimator T for $\tau(\theta)$, which is assumed to be a differentiable function of θ , satisfies the inequality

$$V(T) \geq \frac{[\tau'(\theta)]^2}{E\left[\frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta)\right]^2} \dots (1)$$

Proof We have

$$1 = \int_A \dots \int f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n$$

Differentiating it with respect to θ and using assumption (i), we have

$$0 = \int_A \dots \int \frac{\partial}{\partial \theta} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n$$

$$\text{or } 0 = \int_A \dots \int \frac{\partial}{\partial \theta} [\log f(x_1, \dots, x_n; \theta)] \cdot f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \dots (2)$$

$$\Rightarrow E\left[\frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta)\right] = 0$$

Again T being an unbiased estimator of $\tau(\theta)$,

$$\tau(\theta) = \int_A \dots \int t \cdot f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n = E[T]$$

Differentiating it with respect to θ and using assumption (v) we have

$$\begin{aligned} \tau'(\theta) &= \frac{\partial}{\partial \theta} \tau(\theta) = \int_A \dots \int t \cdot \frac{\partial}{\partial \theta} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \\ &= \int_A \dots \int t \frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta} \cdot f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \dots (3) \end{aligned}$$

Multiplying (2) by $\tau(\theta)$ and subtracting the product from (3), we get

$$\begin{aligned} \tau'(\theta) &= \int_A \dots \int \{t - \tau(\theta)\} \left\{ \frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta} \right\} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \\ &= E\left[\{T - \tau(\theta)\} \left\{ \frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta} \right\}\right] \\ &= \text{Cov.}\left[T, \frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta}\right] \text{ since from (2) } E\left[\frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta}\right] = 0 \end{aligned}$$

Here T has expectation $\tau(\theta)$ and from (2) the expectation of $\frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta}$ is zero. Hence by Cauchy-Schwarz inequality $[\text{Cov.}(X, Y)]^2 \leq V(X) \cdot V(Y)$, we have by virtue of assumption (v)

$$[\tau'(\theta)]^2 \leq V(T) \cdot E\left[\frac{\partial \log f(x_1, \dots, x_n; \theta)}{\partial \theta}\right]^2 \text{ giving the desired result.}$$

Regularity Conditions for Cramer-Rao Inequality (2)

The Cramer-Rao inequality was presented independently and almost simultaneously by Rao, Cramer, Darmois and Fréchet.

Suppose θ is a single parameter varying over the parameter space (H) and that x_1, x_2, \dots, x_n are all continuous variables with joint pdf $f(x_1, x_2, \dots, x_n; \theta)$. We make the following assumptions:

- i) (H) is a non-degenerate open interval on the real line.
- ii) For almost all x_1, x_2, \dots, x_n and all $\theta \in (H)$

$$\frac{\partial}{\partial \theta} f(x_1, x_2, \dots, x_n; \theta)$$

exists, the exceptional, if any, being independent of θ

$$\text{iii) } \frac{\partial}{\partial \theta} \int_A \dots \int f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n = \int_A \dots \int \frac{\partial}{\partial \theta} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n$$

'A' denoting the domain of positive probability density.

$$\text{iv) } \frac{\partial}{\partial \theta} \int_A \dots \int t(x_1, \dots, x_n) f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \\ = \int_A \dots \int t(x_1, \dots, x_n) \frac{\partial}{\partial \theta} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n$$

$$\text{v) } E \left[\frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta) \right]^2 \text{ exists and is positive for every } \theta \in (H).$$

Here, assumption ii) and iv) require that the domain of positive

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Proof: We have

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Differentiating it with respect to θ and using assumption (i), we have

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or $0 = \int_A \dots \int \frac{\partial}{\partial \theta} [\log f(x_1, \dots, x_n; \theta)] \cdot f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \dots (2)$

$$\Rightarrow E\left[\frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta)\right] = 0$$

Again T being an unbiased estimator of $\tau(\theta)$,

$$\tau(\theta) = \int_A \dots \int t \cdot f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n = E[T]$$

Differentiating it with respect to θ and using assumption (iv) we have

$$\tau'(\theta) = \frac{\partial}{\partial \theta} \tau(\theta) = \int_A \dots \int t \cdot \frac{\partial}{\partial \theta} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n \dots (3)$$

Multiplying (2) by $\tau(\theta)$ and subtracting the product from (3), we get

$$\tau'(\theta) = \int_A \dots \int \{t - \tau(\theta)\} \left\{ \frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta) \right\} f(x_1, \dots, x_n; \theta) dx_1 \dots dx_n$$
$$= E\left[\{T - \tau(\theta)\} \left\{ \frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta) \right\}\right]$$

$$= \text{Cov.}\left[T, \frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta)\right] \text{ since from (2) } E\left[\frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta)\right] = 0$$

Here T has expectation $\tau(\theta)$ and from (2) the expectation of $\frac{\partial}{\partial \theta} \log f(x_1, \dots, x_n; \theta)$ is zero. Hence by Cauchy-Schwarz inequality $[\text{cov.}(X, Y)]^2 \leq V(X) \cdot V(Y)$, we have by virtue of assumption (v)

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C B C S. Course

M. Sc. Statistics

Second Semester

Paper-I. Statistical Inference

Unit-II Title.

"CRAMER - RAO INEQUALITY."

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